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## The Neural Basis of Approach-Avoidance Conflict: A Model Based Analysis

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52 **Abstract**

53

54 Approach-avoidance conflict arises when the drives to pursue reward and avoid harm are  
55 incompatible. Previous neuroimaging studies of approach-avoidance conflict have shown large  
56 variability in reported neuroanatomical correlates. These prior studies have generally neglected  
57 to account for potential sources of variability, such as individual differences in choice  
58 preferences and modeling of hemodynamic response during conflict. In the present study, we  
59 controlled for these limitations using a hierarchical Bayesian model (HBM). This enabled us to  
60 measure participant-specific per-trial estimates of conflict during an approach-avoidance task.  
61 We also employed a variable epoch method to identify brain structures specifically sensitive to  
62 conflict. In a sample of 28 human participants, we found that only a limited set of brain  
63 structures (inferior frontal gyrus, right dorsolateral prefrontal cortex and right pre-supplementary  
64 motor area) are specifically correlated with approach-avoidance conflict. These findings suggest  
65 that controlling for previous sources of variability increases the specificity of the  
66 neuroanatomical correlates of approach-avoidance conflict.

67

68 **Significance Statement**

69 Approach-avoidance conflict is implicated in many psychiatric syndromes. Previous fMRI  
70 studies of this important process have potential biases caused by overlooking individual  
71 differences in the evaluation of reward and threat in their analyses. We present a method to  
72 model individual differences in approach-avoidance conflict and demonstrate how to incorporate  
73 this model into fMRI analyses. We found our approach to have greater specificity than previous  
74 studies, which highlights the importance of capturing large variability in participant behavior.

75

76 **Keywords**

77

78 Approach-Avoidance Conflict, fMRI, Hierarchical Bayesian Modeling

## 79 Introduction

80           The drive for self-preservation is fundamental to every living organism. Behavioral  
81 psychologists have long argued that animals evaluate objects and events in their environments  
82 along an appetitive-aversive continuum (Elliot, 2008; Corr, 2013), where animals are motivated  
83 to approach things that sustain them (e.g. rewarding or pleasurable stimuli) and to avoid things  
84 that threaten them (e.g. harmful or painful stimuli). Approach-avoidance conflict arises in  
85 situations where these drives are incompatible, such as when the approach towards reward also  
86 increases the possibility of danger. Approach-avoidance conflict is an important phenomenon as  
87 it is thought to be core to the etiology and maintenance of psychiatric disorders including  
88 depression and anxiety (American Psychiatric Association, 2013).

89           In recent years, many studies have investigated the neural substrates underlying  
90 approach-avoidance conflict using electrophysiology in rodents (Friedman et al., 2015) and non-  
91 human primates (Amemori et al., 2015) and neuroimaging in humans (Talmi et al., 2009; Park et  
92 al., 2011; Bach et al., 2014; Aupperle et al., 2015; O’Neil et al., 2015; Schlund et al., 2016; Loh  
93 et al., 2017). The results of the human neuroimaging literature have implicated a diverse  
94 collection of brain structures in approach-avoidance conflict including cortical structures such as  
95 the anterior cingulate, insula, orbitofrontal cortex, and dorsolateral prefrontal cortex, and  
96 subcortical structures including the amygdala, hippocampus, and striatum. There is considerable  
97 heterogeneity in these findings, however, such that none of the aforementioned brain structures  
98 are consistently identified as being involved in approach-avoidance conflict across these studies.  
99 This naturally prompts the question of where some of the variability might stem from.

100           One possibility is that the heterogeneity reflects variability in approach-avoidance  
101 behavior across participants. Approach-avoidance tendencies are naturally varying across

102 individuals (Carver & White, 1994), such that there are robust individual differences in the  
103 valuation of reward and threat cues. As such, the point of maximal approach-avoidance conflict  
104 is unlikely to be the same across participants. Ignoring these individual differences and averaging  
105 across them, however, has been shown to reduce contrast statistics in fMRI group level analysis  
106 (Ahn et al., 2011). One solution is to explicitly model individual differences in approach-  
107 avoidance conflict, such as with hierarchical Bayesian modeling (Kruschke, 2015), and  
108 incorporate trial-by-trial estimates of approach-avoidance conflict into the fMRI analysis in order  
109 to align participants along a latent evaluation space (O’Doherty et al., 2007; Ahn et al., 2011). In  
110 doing so, we are less likely to average out conflict-related changes in BOLD signal.

111         A second possibility is that the heterogeneity in findings directly reflects variability in  
112 previous modeling of conflict-related changes in BOLD signal. A hallmark feature of approach-  
113 avoidance conflict is prolonged reaction times. Interpreting changes in BOLD signal between  
114 two conditions that also involve differences in response times is challenging, however, due to the  
115 time-on-task effect (Taylor et al., 2014). Because the BOLD signal sums approximately linearly  
116 as a function of stimulation duration (Dale & Buckner, 1997), brain structures not directly  
117 involved in the representation of approach-avoidance conflict may still show increases in BOLD  
118 signal by virtue of prolonged processing of the constitutive elements of conflict (e.g. rewarding  
119 or threatening stimuli). Controlling for response time is necessary then to identify brain  
120 structures that are directly involved in the processing of approach-avoidance conflict (brain  
121 regions that show increased intensity of activity, not just prolonged activity). With the exception  
122 of Talmi et al. (2009), the neuroimaging studies of approach-avoidance conflict cited above do  
123 not document having incorporated response times into their fMRI analyses.

124           In the present study, we investigated the neural signatures of human approach-avoidance  
125 conflict with functional neuroimaging controlling for the issues discussed above. We measured  
126 changes in the fMRI BOLD signal as participants completed an approach-avoidance conflict  
127 task. In the task, participants repeatedly chose between a risky option, returning greater reward at  
128 the risk of potential electrical stimulation, and a safe option, returning a much smaller reward but  
129 no risk of electrical stimulation. Using a novel hierarchical Bayesian model (HBM), we  
130 estimated participants' per-trial approach-avoidance conflict and used these to inform our fMRI  
131 analyses. Moreover, we controlled for the time-on-task effect using the variable epoch method  
132 (Grinband et al., 2008) in order to identify brain structures that showed greater intensity of  
133 activity, rather than prolonged activity, during approach-avoidance conflict. We found that using  
134 these methods increased the specificity of the structures responding to conflict.

## 135 **Methods**

### 136 **Subjects**

137           Thirty-six individuals (13 females, 23 males, age:  $M = 33.94$  yrs,  $SD = 8.80$ ) were  
138 recruited from the Greater Boston area to participate as healthy volunteers in a research program  
139 to develop novel deep brain stimulation (DBS) technologies (Widge et al., 2017). All participants  
140 reported being right-handed and without a current or past diagnosis of a psychiatric or  
141 neurological disorder and were in the normal healthy range for the Mini-International  
142 Neuropsychiatric Interview (MINI; Sheehan et al., 1998). Women were scanned at or near the  
143 ovulation phase of their menstrual cycles (when estradiol is lowest) to minimize potential gender  
144 confounds (Zeidan et al., 2011). The study was approved by the Partners Healthcare System  
145 Human Research Committee, and all participants provided written informed consent prior to

146 enrollment. Participants were paid \$600 for the successful completion of the larger study  
147 protocol.

148         Eight individuals were excluded from analysis: five due to technical complications (see  
149 below), two for missing responses for more than 20% of trials, and one due to corrupted  
150 DICOMs. This resulted in a final sample of 28 participants (10 females, 18 males).

### 151 **Task**

152         We employed a modified version of the Aversion-Reward Conflict (ARC) task (Sierra-  
153 Mercado et al., 2014). During this task, participants make a series of choices between two  
154 options: a safe option or a risky option (Figure 1). Selecting the safe option returns a reward of  
155 \$0.01 and the participant never receives electrical stimulation. In contrast, selecting the risky  
156 option returns a reward between \$0.05 and \$0.95, and the participant receives electrical  
157 stimulation with probability 10%, 50% or 90%, as indicated by a bar in the center of the screen.  
158 This required participants to evaluate their preference for a greater reward with a risk of  
159 electrical stimulation relative to a lesser reward with no risk of electrical stimulation. Participants  
160 were instructed to choose as fast as possible without choosing randomly and were informed that  
161 their choices would be reflected in their final study payment. (In fact, each participant was  
162 compensated with a generous flat payment.) Prior to starting the task, participants were asked to  
163 report back the instructions so that their comprehension could be verified. Next, participants  
164 completed ten practice trials to become accustomed to the timing of the task.

165         This ARC task had three levels of risk: 10%, 50%, and 90% likelihood of electrical  
166 stimulation. Rewards were sampled from all cent values between \$0.05 and \$0.95. Trials were  
167 counterbalanced such that there were an equal number of trials at each risk level, while rewards  
168 were uniformly and equally sampled within each risk level. Each participant completed 108 total

169 trials and the order of trials was kept constant for all the participants. Long inter-trial intervals of  
170  $10.5 \pm 0.875$  seconds separated sequential trials in the task (a slow event-related design). The  
171 duration of the full task was 28.5 minutes.

172       Electrical stimulation was administered to the ankle through a Coulbourn Aversive  
173 Finger Stimulator (Harvard Apparatus, E13-22; maximum level of stimulation = 4.0 mA). The  
174 amperage of electrical stimulation was calibrated individually for each participant prior to  
175 performing the ARC task. Participants experienced increasing levels of stimulation until they  
176 reported reaching a subjective threshold qualified as “highly annoying but not painful”. For five  
177 participants this threshold could not be established because the highest stimulation setting of 4.0  
178 mA was too painful, but penultimate 2.3 mA setting was not experienced as annoying. These  
179 participants did not exhibit behavioral variation (i.e., they always accepted the risky choice) and  
180 consequently these participants were excluded from the analysis.

### 181 **Behavioral Analysis**

182       Our aim was to infer the level of approach-avoidance conflict experienced by each  
183 participant during every trial. We devised a novel HBM that predicts participants’ choices (safe  
184 or risky option) and response times. The decision to model response times was motivated by  
185 well-documented relationship between decision conflict and prolonged response times and prior  
186 demonstrations that including response times in behavioral models improved the accuracy of  
187 single-trial parameter estimation (Prerau et al., 2009; Pederson et al., 2016). The model is  
188 composed of a logistic regression on the choice data and a gamma regression on the response  
189 times. We assume the binary choice responses,  $y \in \{0 = \text{safe choice}, 1 = \text{risky choice}\}$  are drawn  
190 from the Bernoulli distribution:

$$p(y_{ij}|\theta_{ij}) = \theta_{ij}^{y_{ij}}(1 - \theta_{ij})^{1-y_{ij}}$$



191 where  $\theta_{ij}$  is the likelihood-of-take for trial  $i$  and participant  $j$ , and is itself estimated from:

$$\theta_{ij} = \text{logistic}(\beta_{0j} + \sum \beta_{nj} x_{nij})$$

192 Here,  $\beta_{0j}$  is the intercept for participant  $j$ ; the remaining  $\beta_{nj}$  regression coefficients reflect  
 193 the modulatory influence of independent variables,  $X$ , on the baseline likelihood-of-take. In this  
 194 model, there are three independent variables: 50% risk ( $\beta_1$ ), 90% risk ( $\beta_2$ ), and reward ( $\beta_3$ ). The  
 195 50% risk ( $\beta_1$ ) and 90% risk ( $\beta_2$ ), coefficients are binary predictors, whereas reward ( $\beta_3$ ) is a  
 196 continuous predictor that was normalized to have mean = 0 and standard deviation = 1. The  
 197 intercept term,  $\beta_0$ , thus reflects the likelihood of take for 10% risk and \$0.50 reward offer.

198 The continuous response times,  $z$ , are assumed to be drawn from the gamma distribution:

$$p(Z_{ij}|k_j, \mu_{ij}) = \text{Gamma}\left(k_j, \frac{k_j}{\mu_{ij}}\right)$$

199 where  $k_j$  is the shape parameter for participant  $j$  and  $\mu_{ij}$  is the mean of the distribution predicted  
 200 by:

$$\mu_{ij} = \alpha_{0j} + \alpha_{1j} \cdot d_{ij}$$

201 We chose a gamma distribution because it is well-suited for characterizing response times and  
 202 other strictly positive data with a long rightward tail (Yousefi et al., 2015). Here,  $\alpha_{0j}$  was the  
 203 average response time for participant  $j$  and  $\alpha_{1j}$  was the slope term determining how much  
 204 response time increased with conflict. We represent conflict,  $d_{ij}$ , as the inverse of the distance-  
 205 to-decision boundary of trial  $i$  for participant  $j$ , represented as:

$$d_{ij} = 0.25 - (0.5 - \theta_{ij})^2$$

206 This measure,  $d$ , has the shape of an inverted parabola. It is greatest when  $\theta = 0.5$ , or  
 207 when a participant is equally likely to select the safe or risky option. It is smallest when  $\theta = 0.0$   
 208 or  $\theta = 1.0$ , or when a participant is most likely to select the safe or risky option, respectively.

209 Therefore,  $d$  reflects the degree of conflict a participant experienced during the evaluation phase  
210 of a given trial. The model fit then identified the set of parameters that maximized the joint  
211 likelihood of both the choice and response time data due to the relationship between  $\theta$  and  $d$ .

212 As a hierarchical model, each of the participant-level regression parameters defined  
213 above (e.g.  $\alpha_{0j}, \alpha_{1j}, \beta_{0j}, \beta_{1j}, \dots, \beta_{nj}$ ) are drawn a corresponding group-level distribution, centered  
214 at group-level means (e.g.  $\alpha_{0G}, \alpha_{1G}, \beta_{0G}, \beta_{1G}, \dots, \beta_{nG}$ ). Thus, the model simultaneously estimates  
215 group- and participant-level parameters, partially pooling the data so as to minimize the  
216 influence of outliers. Figure 2 presents a detailed diagram of the model which includes the choice  
217 of priors. We assumed Student's t-distribution priors on the choice ( $\beta$ ) regression coefficients to  
218 ensure robust logistic regression (Gelman et al., 2008; Ghosh et al., 2017) using the  
219 recommended degrees of freedom,  $\eta = 5$  (Stan Development Team, 2017).

220 The behavioral model was fit using Hamiltonian Monte Carlo (HMC) sampling in Stan  
221 v2.15 (Carpenter et al., 2017) with four chains of 2000 steps each (1000 burn-in, thinning = 4),  
222 yielding 1000 posterior samples total. The convergence of the chains was computed using the  $\hat{R}$   
223 statistic (Gelman et al., 2014), which measures the degree of variation between chains relative to  
224 the variation within chains. The Stan development team recommends as a rule of thumb that all  
225 parameters have  $\hat{R}$  statistics no greater than 1.1. All parameters in our showed good convergence  
226 ( $\hat{R} \approx 1$ ). Similarly, the number of effective samples approached 1000 for most parameters  
227 indicating that the chains exhibited low autocorrelation. Once fitted, per-trial estimates of  $d$  were  
228 generated by multiplying the observed trial features (risk level and reward value) by the modal  
229 individual-level parameter estimates.

230 **Image Acquisition and Preprocessing**

231 All MRI scans were completed at the Athinoula A. Martinos Center for Biomedical  
232 Imaging. Of the 28 participants included in this analysis, 20 were scanned using a 3T Siemens  
233 Trio scanner and 8 were scanned using a 3T Siemens Prisma scanner (scanner type was entered  
234 as a covariate in the analyses). All participants were scanned with a 32-channel head coil. Foam  
235 cushions were used to restrict head movements. Task images were projected using a rear  
236 projection system and PsychToolbox (V3) stimulus presentation software (Kleiner et al., 2007).

237 For each participant, both structural and functional images were collected. The structural  
238 sequences involved a high-resolution, four-multiecho, T1-weighted, magnetization-prepared,  
239 gradient-echo image (TR = 2510 ms, TE = 1.64 ms, flip angle = 7°, voxel size = 1.0 x 1.0 x 1.0  
240 mm) (van der Kouwe et al., 2008). Functional images were acquired using a multiband SMS-3  
241 T2\*-weighted echo-planar-imaging (EPI) sequence sensitive to blood-oxygen-level dependent  
242 (BOLD) contrast (TR = 1750 ms, TE = 30 ms, flip angle = 75°, voxel size = 2.0 x 2.0 x 2.0 mm,  
243 PAT = GRAPPA, accelerated factor TE = 2). Sixty-three interleaved slices were aligned  
244 perpendicular to the plane intersecting the anterior and posterior commissures, and the whole  
245 brain was imaged (FOV = 220 mm). For the purpose of EPI-dewarping, a fieldmap was also  
246 collected for each participant (63 interleaved slices, TR = 500 ms, TE 1 = 3.41 ms, TE 2 = 5.87  
247 ms, flip angle = 55°, voxel size = 2.0 x 2.0 x 2.0 mm).

248 Anatomical reconstructions of each participant's brain were generated from the T1  
249 structural image using Freesurfer v5.3 (Fischl, 2012). The functional data were first corrected for  
250 slice timing using the Fourier phase shift interpolation from SPM8 and then for B0 using FSL's  
251 *epidewarp* (<http://www.nmr.mgh.harvard.edu/~greve/fbirn/b0/epidewarp.fsl>). FS-FAST v5.3  
252 was used for subsequent preprocessing with their default settings: coregistration with the

253 corresponding Freesurfer anatomical reconstruction; motion correction to the first acquisition  
254 using the AFNI motion correction tool (<http://afni.nimh.nih.gov/afni/>); normalization to  
255 fsaverage/MNI space; and smoothing using 6 mm FWHM kernel.

## 256 **fMRI Modeling and Analysis**

257       Neuroimaging analyses were limited to *a priori* regions of interest in line with the  
258 literature (Talmi et al., 2009; Park et al., 2011; Bach et al., 2014; Aupperle et al., 2015; O’Neil et  
259 al., 2015; Schlund et al., 2016; Loh et al., 2017). Specifically, a cortical mask was constructed  
260 for left and right hemispheres using the Mindboggle atlas (Klein and Tourville, 2012) consisting  
261 of areas encompassing the cingulate cortex, dorsomedial prefrontal cortex, orbitofrontal cortex,  
262 dorsolateral and ventrolateral prefrontal cortex, and insular cortex (Figure 3). Similarly, a  
263 subcortical mask was constructed using the automated subcortical segmentation standard in  
264 Freesurfer (Fischl et al., 2002) consisting of the bilateral striatum (caudate, putamen),  
265 hippocampus, and amygdala.

266       In the first level analysis, we modeled the deliberation phase (time to response) using the  
267 variable epoch method (Grinband et al., 2008). The deliberation phase was modeled using two  
268 sets of boxcar regressors: one control regressor and one parametric modulation regressor. For  
269 both regressors, the boxcar for each trial was scaled in duration according to that trial’s observed  
270 response time. The boxcar for each trial in the parametric modulation term was scaled in  
271 amplitude according to estimated decision conflict ( $d$ ) for that trial. The parametric modulation  
272 boxcars corresponding to trials with missing responses were scaled to zero amplitude.

273       Additionally, several separate control analyses were performed with the same procedure to  
274 determine the effect of 1) using the variable epochs method, 2) using an HBM to model  
275 individual differences and 3) using conflict as the parametric modulator over and above using

276 risk or reward as the parametric modulator. For the first control analysis, fixed epochs were used  
277 instead of variable epochs, where the trial duration was not scaled and instead was uniform; from  
278 the presentation of the first stimulus (the risk bar) to 3.5 seconds after that time when subject  
279 responses were cut off. For the second control analysis, an equivalent non-hierarchical model  
280 was used (i.e. estimating only group parameters, excluding participant-level parameters) to  
281 model the conflict parametric modulator. For the third set of control analyses, risk and reward  
282 were used, in separate analyses, to parametrically modulate the deliberation-phase regressor  
283 instead of conflict, and, in another separate analysis, risk, reward and conflict were all used as  
284 parametric modulators with three parametrically modulated deliberation-phase regressors in the  
285 same first-level analysis. All regressors were convolved with the SPM hemodynamic response  
286 function. All estimated regression coefficients in first level analysis were converted to percent  
287 signal change (PSC; Pernet, 2014).

288         The fMRI data were preprocessed using a high-pass filter, nuisance regressors and  
289 motion scrubbing. A discrete cosine transform basis set was added to high-pass filter the data at  
290 0.01 Hz. The six possible directions of motion were incorporated into the first-level analyses  
291 (after being demeaned, detrended, and orthogonalized) as nuisance regressors. Finally, motion  
292 scrubbing was used to mitigate the impact of high-motion acquisitions on the data (Siegel et al.,  
293 2014). Volumes for which the calculated framewise displacement (Power et al., 2012) exceeded  
294 0.9 mm were excluded from analyses, and the first four acquisitions were discarded.

295         In the second level analysis, the beta coefficients estimated for each participant were  
296 submitted to a weighted least squares (WLS) regression where F-contrasts were computed for the  
297 control and parametrically modulated regressors. Scanner type (Trio vs. Prisma) was entered as a  
298 secondary nuisance regressor. Five thousand permutations of the WLS model were also

299 computed following the Freedman–Lane procedure (Winkler et al., 2014). Every statistical map,  
300 both observed and permuted, was submitted to threshold-free cluster enhancement (Smith and  
301 Nichols, 2009) using the recommended parameters ( $H = 2$ ,  $E = 0.5$ ,  $\text{step} = 0.1$ ). Finally, the  
302 permutation maps were used to compute family-wise error (FWE) corrections ( $\alpha = 0.05$ ) for each  
303 voxel (Winkler et al., 2014). Any resulting clusters were discarded if they covered less than 100  
304  $\text{mm}^2$  on the surface or fewer than 20 contiguous voxels of the volume.

### 305 **Code Accessibility**

306 All data and analysis scripts are available online at [openneuro.org/datasets/ds001814](https://openneuro.org/datasets/ds001814) and  
307 [github.com/mghneurotherapeutics/DARPA-ARC](https://github.com/mghneurotherapeutics/DARPA-ARC) respectively. The data and scripts are freely  
308 available at these locations with instructions for access and suggested citation included.

## 309 **Results**

### 310 **Behavioral Results**

311 Participants exhibited the expected response trends for the ARC task: greater risk of  
312 electrical stimulation decreased on average the likelihood of selecting the risky option, whereas  
313 increasing reward increased the likelihood of selecting the risky option (Figure 3). The 95%  
314 highest density intervals (HDIs) of the posterior distribution for the group-level parameters  
315 showed decreases in risky-choice taking for the 50% ( $\beta_1 = -1.922$ , 95% HDI: [-2.606, -1.139])  
316 and 90% risk ( $\beta_2 = -4.180$ , 95% HDI: [-5.273, -3.257]) conditions. In contrast, increases in risky-  
317 choice taking were observed in response to increasing reward ( $\beta_3 = 10.652$ , 95% HDI: [8.239,  
318 12.887]). Thus, risk biased choice behavior towards avoidance (i.e. selecting the safe option), and  
319 reward biased choice behavior towards approach (i.e. selecting the risky option), indicating that  
320 the ARC task elicited the intended behavioral effects.

321           At the subject level, the 95% HDIs of the posterior estimates for the 50% risk ( $\beta_1$ ) and  
322 90% risk ( $\beta_2$ ) coefficients were strictly negative for 19/28 participants and 24/28 participants,  
323 respectively. The 95% HDIs for the reward coefficients ( $\beta_3$ ) were strictly positive for 27/28  
324 participants. No participants exhibited an increase in choice preference for the risky option with  
325 increasing risk and no participants exhibited an increase choice preference for the safe option  
326 with increasing reward. In summary, all the participants had response trends that matched our  
327 expectations for the ARC task, and most participants' behavior was modulated by both risk and  
328 reward.

329           For the response time component of our HBM, we found that approach-avoidance  
330 conflict was positively correlated with response times (Figure 4B). At the group-level, the 95%  
331 HDI of the posterior distribution on the conflict-RT slope parameter was strictly positive ( $\alpha_1 =$   
332 0.456, 95% HDI: [0.388, 0.528]). The model estimated an average increase in response times of  
333 0.456 seconds at maximal conflict. Thus, the ARC task was also successful in eliciting this  
334 hallmark behavioral signature of increased response times during approach-avoidance conflict.

335           It is important to note we observed considerable variability in the choice preferences of  
336 our participants (Figure 5). The most approach-biased participant selected the risky option on  
337 almost all trials (93%), whereas in contrast the most avoidance-biased participant selected the  
338 safe option on almost all trials (16%). This strongly demonstrates the notion that the points of  
339 maximal approach-avoidance conflict are unlikely to be the same across participants and  
340 reinforces the need for methods like HBMs that explicitly take into consideration these large  
341 individual differences.

342           Importantly, posterior predictive checks showed that our model accurately captured  
343 participants' choice behavior (Figure 5). The root-mean-square error between predicted and

344 observed average risky choice was 0.023. To assess the possibility of model overfitting, we  
345 compared the widely applicable information criterion (WAIC; Watanabe, 2013) of our HBM to  
346 an equivalent non-hierarchical model (i.e. estimating only group parameters, excluding  
347 participant-level parameters). WAIC scores are reported here on deviance scale where lower  
348 scores denote greater fitness. The hierarchical model (WAIC = 1319.4) was strongly preferred to  
349 its non-hierarchical equivalent (WAIC = 2611.5) despite its greater complexity. We also  
350 compared our hierarchical model to a secondary hierarchical model that included interaction  
351 terms between risk and reward. This model performed slightly worse than the main effects-only  
352 model (WAIC = 1320.8). As such, we proceeded with the more parsimonious model with main  
353 effects only for fMRI analysis.

354         In summary, the ARC task successfully elicited approach, avoidance, and approach-  
355 avoidance conflict behaviors from all participants. Specifically, participants were (1) more likely  
356 to select the risky option with increasing reward; (2) more likely to select the safe option with  
357 increasing risk of electrical stimulation; and (3) slower to respond with increased approach-  
358 avoidance conflict. Moreover, participants exhibited large individual differences in their choice  
359 preferences, which were accurately captured by our HBM. It is worth reiterating that ignoring  
360 these differences can reduce contrast effects in fMRI analysis by averaging over the neural  
361 correlates of dissimilar cognitive processes (Ahn et al., 2011).

## 362 **Imaging Results**

363         For the control regressor (i.e. measuring the average BOLD signal change during the  
364 deliberation phase, without modulation by conflict), we found activations within the *a priori*  
365 cortical and subcortical regions of interest (Figure 6) that were selected based on prior literature  
366 (see the fMRI Modeling and Analysis section). Peak voxels and their corresponding statistics are



367 reported in Table 1. Large, significant BOLD signal increase was observed in bilateral dorsal  
368 anterior cingulate and dorsomedial prefrontal cortex (dACC/dmPFC; BA 32), midcingulate  
369 cortex (BA 23/24), pre-supplementary motor area (pre-SMA; BA 6), anterior insula (BA 13), and  
370 dorsolateral prefrontal cortex (dlPFC; BA 46). Among subcortical structures, the control  
371 deliberation regressor was positively correlated with BOLD signal activation in bilateral dorsal  
372 hippocampus and striatum (caudate, putamen). Smaller, significant activations were also  
373 detected in the right lateral orbitofrontal cortex (OFC; BA 11) and right putamen. These results  
374 corroborate the distributed network of neural structures previously reported to be recruited  
375 during approach-avoidance conflict tasks (Talmi et al. 2009; Park et al., Bach et al., 2014;  
376 Aupperle et al., 2015; O'Neil et al., 2015; Schlund et al., 2016; Loh et al., 2017).

377         Significant change in BOLD signal for approach-avoidance conflict regressor was  
378 observed in a much more restricted set of structures (Figure 7.). Approach-avoidance conflict  
379 was positively correlated with BOLD signal activation only in bilateral rostral inferior frontal  
380 gyrus (IFG; pars orbitalis; BA 47), right dlPFC (BA 46), and right dmPFC/pre-SMA (BA 32).  
381 No significant positive activations were detected in subcortical structures, and no negative  
382 activations were detected in any *a priori* region of interest. In contrast to the aforementioned  
383 previous literature, our results suggest that only a select set of cortical structures tracked  
384 approach-avoidance conflict. Interestingly, our analysis revealed conflict representations in the  
385 right IFG, a structure previously unreported in the approach-avoidance conflict literature.

386         The control analyses showed the difference between these results and results from  
387 analyses with fixed epochs, averaging across subjects and using a simpler risk or reward only  
388 model. As shown in Figure 8, using fixed epochs caused smaller, more widespread, positive  
389 activations encompassing bilateral striatum and left insula in addition to the structures activated

390 in the main, variable epochs analysis. The non-hierarchical Bayesian model (used in combination  
391 with variable epochs) had no significant activations that correlated with the conflict regressor.  
392 Using risk and reward as regressors in a model with only the non-parametrically modulated,  
393 control deliberation regressor and risk or the control deliberation regressor and reward also  
394 yielded almost no significant activations with the exception of a small, negative activation  
395 correlated with reward in a small area of right dlPFC and lateral OFC. When the risk and reward  
396 regressors were modeled in combination with conflict, not only were there no significant  
397 activations for the risk and reward regressors, but the significant activations for the conflict  
398 regressor was suppressed.

## 399 Discussion

400 In this study, we investigated the neural basis of human approach-avoidance conflict  
401 while accounting for two possible sources of heterogeneity in the literature; individual approach-  
402 avoidance variability and time-on-task. Using hierarchical Bayesian modeling, we controlled for  
403 individual differences in approach-avoidance preference by comparing participants' fMRI data  
404 according to each participant's relative points of maximal approach-avoidance conflict. Using  
405 the variable epochs method in our fMRI analyses, we also controlled for the time-on-task effect.  
406 Thus, we were able to differentiate brain structures strictly sensitive to approach-avoidance  
407 conflict from those representing information correlated with deliberation more generally. The  
408 present findings corroborate previous reports of the anatomical correlates of approach-avoidance  
409 behavior by our finding that BOLD signal increased during deliberation across a broad network  
410 of cortical and subcortical brain structures (dACC/dmPFC, pre-SMA, dlPFC, OFC, insula,  
411 striatum, hippocampus). Importantly, the current findings deviate from the previous literature  
412 insofar that our controlled analyses found conflict-related changes in BOLD signal only in a

413 select set of structures (i.e. IFG, dlPFC, and pre-SMA). Collectively, the current findings suggest  
414 the importance of careful methodology in isolating the neuroanatomical correlates of latent  
415 psychological states such as approach-avoidance conflict.

416         To examine the effect of using an HBM, we compared these results to results from the  
417 non-hierarchical Bayesian model analysis. The HMB methodology was clearly warranted by the  
418 large differences observed between the approach-avoidance behavior of different participants as  
419 shown in Figure 5 and described in the Behavioral Results section. The need for this  
420 methodology was confirmed by the suppression of any significant areas of activation when a  
421 non-hierarchical Bayesian model was used. Thus, accounting for individual differences with an  
422 HBM resulted increased group-level fMRI contrast statistics, consistent with previous findings  
423 (Ahn et al., 2011).

424         Another important difference between the present findings and past studies is our use of  
425 the variable epochs method (Grinband et al., 2008), which we included so as to control for the  
426 time-on-task effect and minimize the risk of mismodeling the hemodynamic response. By  
427 controlling for time-on-task, our analysis was explicitly interested in identifying brain structures  
428 that show an increase in the BOLD signal due to an increase in the intensity, not duration, of the  
429 activity of the underlying neural populations. One natural question is whether approach-  
430 avoidance conflict is more accurately modeled as the prolonged, but not increased, engagement  
431 of brain structures. One problem with this view, as noted above, is that this makes it difficult to  
432 disentangle conflict-specific signals from other correlated but unrelated signals (e.g. processing  
433 of reward or threat stimuli). As such, we opted to use a more conservative definition of  
434 approach-avoidance conflict (increase in amplitude of BOLD signal, above and beyond that  
435 expected from prolonged engagement, as measured by our parametric modulation regressor).

436 The conservativeness of this variable epochs method compared to fixed epochs was confirmed in  
437 the control analyses shown in Figure 8, where areas with significant activation in the variable  
438 epochs analysis were found to be a subset of areas with significant activation for the fixed epochs  
439 analysis. Thus, our analysis was conservative but well suited to identify regions specifically  
440 implicated in the processing of approach-avoidance conflict.

441 To control for whether our results relate to approach-avoidance and not some simpler  
442 approach or avoidance alternative mechanism, we ran three different analyses 1) with risk as the  
443 parametrically modulated regressor, 2) with reward as the parametrically modulated regressor  
444 and 3) with three parametrically modulated regressors for risk, reward and conflict. The first two  
445 analyses showed that risk or reward alone are not capable of explaining the regions of conflict  
446 that had significant activations correlated with conflict (Figure 7); as described in the results  
447 section, these analyses had almost no areas of significant activation. In the third analysis, the  
448 suppression of significant conflict activations (described in the results section) suggested that  
449 including risk and reward in the same model as conflict caused the variance to be split between  
450 all three variables' explanatory power. Reward and risk are approach and avoidance stimuli,  
451 respectively, so by definition these stimuli covary strongly with the approach-avoidance  
452 measure—conflict. This control analysis therefore confirms that the explanatory power of  
453 conflict is dependent on risk and reward and also shows that including regressors with high  
454 covariance can cause a false-negative result.

455 Another point worth noting is that our analysis assumes only linear changes in the BOLD  
456 response to conflict. The variable epochs method utilized here is insensitive to any nonlinear  
457 changes in the BOLD signal that may arise as a function of response time, raising the possibility  
458 of remaining biases in the present results. Interestingly, in a finite impulse response analysis of

459 the hemodynamic response during prolonged response times, Yarkoni et al. (2009) found that  
460 structures in the prefrontal cortex were better described by increases in the amplitude of  
461 hemodynamic response but not by changes in its shape. These findings suggest that not using a  
462 finite impulse response analysis did not bias the hemodynamic response in this present analysis,  
463 but further studies are necessary to answer this question more definitively.

464         There were additional discrepancies between the present study and previous studies on  
465 approach-avoidance tasks. Though positive BOLD activation was detected during deliberation in  
466 the right orbitofrontal cortex (OFC), the effect was considerably smaller than previously reported  
467 findings (Talmi et al., 2009; Schlund et al., 2016). This may reflect signal-to-noise ratio issues  
468 particular to surface-based analysis of the OFC (Stenger, 2006). Additionally, in contrast to  
469 Schlund et al. (2011) and Aupperle et al. (2015), amygdala activation was not detected during  
470 deliberation. In this study, suboptimal calibration of the stimulation amperage likely diminished  
471 participants' perception of threat from the stimulation and consequently their amygdala  
472 activation. Finally, the bilateral hippocampus activations detected during deliberation were  
473 located dorsally, rather than anteriorly/ventrally as have been previously reported in literature on  
474 threat processing (Bach et al., 2014). The dorsal hippocampus has been associated with cognition  
475 and planning (Fanselow and Dong, 2010), so these activations could reflect participants'  
476 processing of the conditional structure of the ARC task (e.g. "if *safe* is chosen, then 0% chance  
477 of electrical stimulation; if *risky* is chosen, then X% chance of electrical stimulation").

478         This study had several limitations. Due to the equipment issues described above, as well  
479 as the use of non-adaptive rewards, we were unable to calibrate the reward and risk of the ARC  
480 task according to each participant's choice preferences. This may be one reason why we  
481 observed an approach-bias on average. This also means that the present study undersampled

482 trials at or near the points of participants' maximal approach-avoidance conflict. A consequence  
483 of this undersampling is that many of the high conflict decisions participants made in this task  
484 occurred during high risk trials, making it harder to divorce conflict from risk. Future approach-  
485 avoidance conflict experiments should consider incorporating adaptive design optimization  
486 (Myung et al., 2013) in order to titrate the levels of rewarding and threatening stimuli according  
487 to future participants' choices preferences to minimize the influence of these potential biases.

488         Finally, it is worth noting that the set of structures we found correlated with approach-  
489 avoidance conflict (i.e. IFG, dlPFC and pre-SMA) share overlap with the putative response  
490 inhibition network (Aron et al., 2004, 2014; Aron and Poldrack, 2006). One interpretation of the  
491 present results is that approach-avoidance conflict is another process requiring response  
492 inhibition, wherein the IFG inhibits prepotent motor responses in order to facilitate prolonged  
493 evidence accumulation during difficult choices. This interpretation is consistent with the  
494 increased response times observed in the present experiment. The possible role of the inhibition  
495 network during approach-avoidance conflict points to a clear direction for future studies;  
496 investigating whether the putative response inhibition network works to signal response conflict  
497 to other brain structures, such as through the hyperdirect pathway to the basal ganglia (Frank et  
498 al., 2015). Alternately, these structures may be involved in the resolution of approach-avoidance  
499 conflict, such as by biasing choice towards approach or avoidance. In either case, the framework  
500 that this study presents for the consideration of individual-level behavioral variation and the  
501 time-on-task effect would likely lead to benefits in specificity and accuracy of future studies  
502 investigating similar cognitive processes.

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623 Legends

624

625 **Figure 1: The Aversion-Reward Conflict (ARC) task.** Participants are presented with a safe  
626 choice (blue) and a risky choice (orange). The safe choice pays a guaranteed small reward  
627 (\$0.01) and no aversive stimulation. The risky choice pays a guaranteed larger reward (\$0.05 -  
628 \$0.95), and a probability of stimulation as indicated by the centered white bar. Participants  
629 decide whether to accept a higher payout at risk of aversive stimulation.

630 Figure Contributions: Darin Dougherty, Thilo Deckersbach, Alik Widge and Samuel Zorowitz  
631 designed the task, Sam Zorowitz created the figure.

632

633 **Figure 2: A Kruschke-style diagram of the hierarchical model.** The “~” symbol indicates  
634 stochastic dependency, whereas the “=” symbol indicates a deterministic dependency. Ellipses  
635 indicate the indices over which the dependency applies. The parameter of most interest is  $d$ , the  
636 inverse distance-to-decision-boundary, which measures the estimated conflict experienced on a  
637 given trial.

638 Figure Contributions: Samuel Zorowitz created the model.

639

640 **Figure 3: A priori cortical regions of interest.** Regions (Freesurfer labels) were selected from  
641 the Mindboggle atlas ([mindboggle.info/data.html](http://mindboggle.info/data.html)) based on the diffuse locations of activations  
642 previously reported in the approach-avoidance decision making literature.

643 Figure Contributions: Samuel Zorowitz chose the regions of interest based on prior literature and  
644 created the figure.

645

646 **Figure 4: Group-level behavior results.** (A) The estimated likelihood of choosing the risky  
647 option for each risk level and across rewards. The model estimated decreases in risky decision  
648 making at both 50% risk ( $\beta_1 = -1.922$ , 95% HDI: [-2.606, -1.139]) and 90% risk ( $\beta_2 = -4.180$ ,  
649 95% HDI: [-5.273, -3.257]). In contrast, the model estimated increases in risky decision making  
650 in response to increasing reward ( $\beta_3 = 10.652$ , 95% HDI: [8.239, 12.887]). (B) The estimated  
651 linear component of deliberation time as a function of decision conflict,  $d$ . The model estimated  
652 an increase in deliberation time with decision conflict ( $\alpha_1 = 0.456$ , 95% HDI: [0.388, 0.528]).  
653 Shaded regions denote the 95% highest density interval (HDI).

654 Figure Contributions: Samuel Zorowitz, Katherine Link and Alexander Rockhill performed the  
655 behavioral analysis.

656

657 **Figure 5: Individual differences in behavior.** Participants in the ARC task exhibited large  
658 individual differences in behavior. (A) Participants varied in their approach-avoidance  
659 preferences (though the majority were approach-biased). (B) Participants varied in the extent to  
660 which their deliberation increased in response to decision conflict (but all participants showed  
661 increased response times during conflict). Each point represents one participant. The horizontal  
662 axis denotes the observed behavior (proportion of risky choices, A; response time increases, B),  
663 and the vertical axis denotes the model predicted behavior. Proximity to the diagonal indicates  
664 goodness of fit.

665 Figure Contributions: Samuel Zorowitz, Katherine Link and Alexander Rockhill performed the  
666 behavioral analysis.

667

668 **Figure 6: Percent signal change (PSC) during deliberation.** The control regressor measures  
669 changes in the BOLD signal during deliberation (independent of approach-avoidance conflict).  
670 Positive activation was found in cortical and subcortical regions including the lateral and medial  
671 prefrontal cortex, striatum, and hippocampus. All voxels corrected for multiple comparisons  
672 through 5000-iteration permutation testing and voxel-wise FWE corrections ( $\alpha = 0.05$ ).  
673 Abbreviations: LH = left hemisphere; RH = right hemisphere.  
674 Figure Contributions: Samuel Zorowitz and Alexander Rockhill performed the fMRI analysis;  
675 Samuel Zorowitz, Alexander Rockhill and Kristen Kellard collected the data.

676  
677 **Figure 7: Percent signal change (PSC) during conflict.** The parametric modulation regressor  
678 measures changes in BOLD signal during deliberation as a function of approach-avoidance  
679 conflict. Positive activation was detected only in bilateral inferior frontal gyrus (IFG), and right  
680 dorsolateral prefrontal cortex (dlPFC) and pre-supplementary motor area (pre-SMA). All voxels  
681 corrected for multiple comparisons through 5000-iteration permutation testing and voxel-wise  
682 FWE corrections ( $\alpha = 0.05$ ).  
683 Figure Contributions: Samuel Zorowitz and Alexander Rockhill performed the fMRI analysis;  
684 Samuel Zorowitz, Alexander Rockhill and Kristen Kellard collected the data.

685  
686 **Figure 8: Percent signal change (PSC) during conflict for the fixed epochs analysis.** In this  
687 case, epochs were made from the first stimulus presentation to the end of the response period  
688 instead ending when the subject responded for each particular trial. More widespread, less  
689 specific, smaller, positive activation was detected in the same structures as Figure 7 with the  
690 addition of activation in bilateral striatum, left insula as well as greater activation in bilateral  
691 dlPFC. All voxels corrected for multiple comparisons through 5000-iteration permutation testing  
692 and voxel-wise FWE corrections ( $\alpha = 0.05$ ).  
693 Figure Contributions: Alexander Rockhill performed the fMRI analysis; Samuel Zorowitz,  
694 Alexander Rockhill and Kristen Kellard collected the data.

695  
696 **Table 1: Coordinates and statistics of peak BOLD activations.** The reported statistics are the  
697 percent signal change (PSC) and weighted-least squares (WLS) contrast against baseline ( $F$ )  
698 statistic. The first set of results reflect the unmodulated deliberation and the second set reflect the  
699 contrast between deliberation parametrically modulated by conflict and unmodulated  
700 deliberation. All coordinates reported in the Montreal Neurological Institute (MNI) space and  
701 reflect the peak of activation. All voxel statistics were corrected for multiple comparisons  
702 through 5000-iteration permutation testing and voxel-wise FWE corrections ( $\alpha = 0.05$ ).  
703 Abbreviations: LH = left hemisphere; RH = right hemisphere; dACC = dorsal anterior cingulate  
704 cortex; dmPFC = dorsomedial prefrontal cortex; MCC = midcingulate cortex; DLPFC =  
705 dorsolateral prefrontal cortex; OFC = orbitofrontal cortex; pre-SMA = pre-supplementary motor  
706 cortex; IFG = inferior frontal gyrus.  
707 Figure Contributions: Samuel Zorowitz and Alexander Rockhill performed the fMRI analysis;  
708 Samuel Zorowitz, Alexander Rockhill and Kristen Kellard collected the data.

709 Tables

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<b>Deliberation Phase (Control)</b>					
<b>ROI</b>	<b>X</b>	<b>Y</b>	<b>Z</b>	<b>PSC</b>	<b>F</b>
dACC/dmPFC: LH	-12	22	36	0.08	352.92
RH	7	15	24	0.09	462.05
MCC: LH	-7	-22	29	0.15	328.72
RH	7	-15	31	0.18	529.27
pre-SMA: LH	-9	7	51	0.10	419.56
RH	10	14	47	0.10	373.34
dIPFC: LH	-36	9	24	0.12	223.96
RH	36	18	25	0.11	312.87
Anterior Insula: LH	-31	27	9	0.2	351.91
RH	31	27	8	0.16	413.00
Lateral OFC: RH	13	38	-24	0.07	95.60
pre-Motor: LH	-37	-2	43	0.14	291.16
RH	36	-3	44	0.14	333.73
Caudate: LH	-10	7	3	0.07	28.42
RH	10	11	5	0.06	25.16
Putamen: LH	-20	5	1	0.05	29.04
RH	34	-7	-7	0.04	22.15
Hippocampus: LH	-14	-39	-3	0.09	34.95
RH	14	-39	-1	0.10	34.47

<b>Deliberation Phase (Conflict)</b>					
IFG: LH	-39	45	7	0.05	56.53
RH	42	45	-6	0.05	55.10
dlPFC: RH	42	27	31	0.04	68.02
pre-SMA: RH	9	27	46	0.04	59.55

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