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Lingual Gyrus surface area is associated with anxiety-depression severity in young adults: a genetic clustering approach

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1 **Lingual Gyrus surface area is associated with anxiety-depression severity**
2 **in young adults: a genetic clustering approach**

3

4 **Abbreviated title:** Anxiety-depression and lingual gyrus surface area

5

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32
33 This manuscript consists of 38 pages, 4 figures, 3 tables.

34 The abstract is 224 words long, the significance statement 111, the introduction
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36
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52

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67

68

69 Abstract

70 Here we aimed to identify cortical endophenotypes for anxiety/depression. Our
71 data-driven approach used vertex-wise genetic correlations (estimated from a
72 twin sample: 157 monozygotic and 194 dizygotic twin pairs) to parcellate
73 cortical thickness and surface area into genetically homogeneous regions (Chen
74 et al., 2013). In an overlapping twin and sibling sample (N=834; aged 15-29, 66%
75 female), in those with anxiety-depression SPHERE scores (Hickie et al., 2001)
76 above median we found a reduction of surface area in an occipito-temporal
77 cluster, which comprised part of the right lingual, fusiform and parahippocampal
78 gyrii. A similar reduction was observed in the Human Connectome Project
79 sample (N=890, age 22-37, 56.5% female) in those with Adult Self Report DSM-
80 oriented scores (Achenbach et al., 2005) in the 25-95% quantiles. A post-hoc
81 vertex-wise analysis identified the right lingual and, to a lesser extent the
82 fusiform gyrus. Overall, the surface reduction explained by the anxiety-
83 depression scores was modest ($r=-0.10$, 3rd order spline, and $r=-0.040$, 1st order
84 spline in the HCP). The discordant results in the top 5% of the anxiety-
85 depression scores may be explained by differences in recruitment between the
86 studies and especially medication screening as anti-depressants may increase
87 the lingual gyrus volume (Jung et al., 2014). However, we could not conclude
88 whether this cortical region was an endophenotype for anxiety-depression as the
89 genetic correlations did not reach significance, which we attribute to the modest
90 effect size (post-hoc statistical power<10%).

91

92 Significance Statement

93 Endophenotypes may help shed light on the aetiology, cognitive impairment and
94 genetics of psychiatric disorders. Here, we report a non-linear negative
95 association between anxiety-depression and smaller surface area of the occipito-
96 temporal region, which comprises most of the right lingual and fusiform gyri
97 (N=834). This cluster was defined by applying a fuzzy clustering algorithm to a
98 matrix of vertex-wise genetic correlations among cortical surface measures. We
99 replicated this association in an independent sample from the Human
100 Connectome Project (N=890). We could not confirm the presence of a genetic
101 correlation as the effect size of the association was modest ($r=-0.10$).

102

103 Introduction

104 Endophenotypes may be useful in decomposing or deconstructing a
105 psychiatric disorder, which should result in more successful genetic analyses,
106 redefinition of diagnosis, improved studies of the course of illness and
107 development of pertinent animal models (Gottesman and Gould, 2003).
108 According to a widely accepted definition, an endophenotype needs to satisfy 4
109 core criteria (Gottesman and Gould, 2003): it should be heritable, genetically and
110 phenotypically associated with the illness in the population and state-
111 independent (not a consequence of the disorder). There should also be a genetic
112 correlation between the disease and endophenotype: unaffected family members
113 of affected individuals should also exhibit the endophenotype, to some extent, as
114 they have a higher genetic load for the disease than people who are not closely
115 related to affected individuals. Reliability of the endophenotype is sometimes

116 included as part of the definition (Goldstein and Klein, 2014) and is implied by
117 the trait heritability (Dohm, 2002).

118 The endophenotype concept is popular (Glahn et al., 2014), but only a few
119 endophenotypes have been identified for affective disorders (Glahn et al., 2012;
120 Goldstein and Klein, 2014; Bas-Hoogendam et al., 2016). These include:
121 neuroticism (Goldstein and Klein, 2014), personal life events (Boardman et al.,
122 2011) and perceived social support (Kendler and Baker, 2007; Matthews et al.,
123 2016). Magnetic resonance imaging (MRI) brain measures may be promising
124 endophenotypes, as they are objective (not self-reported), reliable (Liem et al.,
125 2015) and heritable (Strike et al., 2015; Whelan et al., 2016). There has been an
126 explosion of proposed brain endophenotypes for depression, and most
127 publications report unreplicated phenotypic associations based on small sample
128 sizes (see (Savitz and Drevets, 2009; Hasler and Northoff, 2011; Zhang et al.,
129 2013) for reviews). To overcome publication bias and propose robust candidate
130 endophenotypes, the Major Depressive Disorder (MDD) group of the ENIGMA
131 consortium (Thompson et al., 2014) has conducted large-scale case-control
132 comparisons, which found several structural brain phenotypes consistently
133 associated with MDD (Schmaal et al., 2015; Schmaal et al., 2016), but somewhat
134 conflicting cortical signatures of depression when stratifying the analysis by age
135 (cut-off 21 yrs.). In addition, another large voxel-wise meta-analysis of cortical
136 grey matter density reported a profile of volume changes associated with MDD
137 (Arnone et al., 2016). However, more work is needed to determine genetic
138 relationships between depression and structural brain measures and to replicate
139 the reported associations.

140 Here, we aimed to identify cortical thickness (CT) or surface area (SA)
141 endophenotypes for depression and anxiety, which share most of their genetic
142 risk (Kendler et al., 1987; Hettema, 2008). We considered both CT and SA as
143 there is evidence they are influenced by specific environmental and genetic
144 factors, and thus are preferred to combined volumetric measures as they gather
145 more information (Panizzon et al., 2009; Winkler et al., 2010). We performed an
146 exploratory data driven analysis in a large, genetically informative young adult
147 twin sample, followed by a replication in an independent cohort of similar age.
148 The large cohorts allowed us to investigate the presence of non-linear
149 relationships between cortical structures and anxiety-depression level.

150 We used twin modelling (Martin and Eaves, 1977; Neale and Cardon,
151 1992) to perform a genetic parcellation of the cortex (Chen et al., 2013), aiming
152 to reduce the data dimension, while ensuring genetic homogeneity of the brain
153 phenotypes studied. We controlled for mean SA/CT to adjust for the confounding
154 effect of overall brain or body size. To measure anxiety-depression, we used a
155 validated continuous measure of anxiety-depression (SPHERE: Somatic and
156 Psychological Health Report) (Hickie et al., 2001) collected before the MRI scan,
157 rather than a DSM-IV lifetime diagnosis (with disease onset not always preceding
158 imaging). We could then get around the question of state-independence (even if
159 a longitudinal design is needed to confirm the direction of causation) and
160 maximised statistical power by providing a larger sample with greater variance
161 across the population. In the replication sample, we used validated Achenbach
162 scales (Achenbach et al., 2005; Achenbach, 2009) collected at the time of
163 imaging. Furthermore, we aimed to characterise sources of covariation between
164 anxiety-depression and cortical measures using twin modelling.

165

166 Materials and Methods

167 Exploratory Analysis

168 Participants

169 We analysed data from a total of 834 twins and siblings from the Brisbane
170 Longitudinal Twin Study (BLTS)(Wright and Martin, 2004; Gillespie et al., 2013)
171 who were assessed for symptoms of anxiety-depression and who had undergone
172 brain imaging. Anxiety-Depression was assessed at mean age 17 (SD=2.2, range
173 10-24) with imaging completed 4.4 years (SD=2.1, range 0.6 to 9.5 years) later, at
174 a mean age of 21 (SD=3.2, range 15-29). The sample (66% female) comprised 23
175 twin-sibling trios (9 MZ, 14 DZ), 275 complete pairs (101 MZ, 142 DZ and 32
176 pairs of siblings treated as DZ) and 214 singletons (Table 1). Zygosity of twin
177 pairs was initially determined from DNA using a commercial kit (AmpFISTR
178 Profiler Plus Amplification Kit, ABI) and was later confirmed by genome-wide
179 single nucleotide polymorphism genotyping (Illumina 610K chip).

180

181 Brain Imaging

182 Imaging was conducted using a 4T Bruker Medspec whole-body MRI on
183 1,161 twins and their siblings aged 15 to 30 (62% female), as part of the QTIM
184 study of brain structures and function (de Zubicaray et al., 2008; Blokland, 2014;
185 Renteria et al., 2014). Participants were all right-handed and were screened for
186 prior mental health diagnoses and anti-depressant use, as well as neurological
187 disorders and loss of consciousness. Structural T1-weighted 3D images were
188 acquired with the following parameters: TR=1500ms, TE=3.35ms, TI=700ms,
189 240mm FOV, 0.9mm slice thickness, 256 or 240 slices depending on acquisition

190 orientation: 86% coronal (256 slices), 14% sagittal (240 slices). The raw T1-
191 weighted images were corrected for intensity inhomogeneity with SPM8
192 (Flandin and Friston, 2008; Wellcome Department of Imaging Neuroscience,
193 2009). Cortical surfaces were reconstructed using FreeSurfer (v5.3) (Fischl and
194 Dale, 2000; Fischl, 2012), then resampled into a common space, smoothed (N =
195 2819 iterations, nearest neighbour) and down-sampled, with SA and CT
196 measured at each surface vertex (2562 per hemisphere). Whole brain total SA
197 and mean CT were additionally extracted. : From the total imaging sample
198 (N=1,161), we excluded 11% of participants due to neuroanatomical
199 abnormalities (n=48), excessive head motion during scanning (n=16), or poor
200 quality FreeSurfer cortical surface reconstructions (n=57, using the ENIGMA
201 quality checking procedure that consists of manual QC assisted by automatic
202 detection of outliers (enigma.ini.usc.edu)).

203

204 Assessment of anxiety-depression

205 The anxiety-depression score was calculated using the SPHERE
206 questionnaire (Hickie et al., 2001), which has been administered to 3,312 twins
207 and siblings across 1 to 4 waves of the BLTS. When several SPHERE scores were
208 available, we selected the closest to the time of scan. We previously showed that
209 this score is moderately heritable (Hansell et al., 2012) and has good
210 psychometric properties (stochastic ordering on the sum score, 3 months' test-
211 retest, internal consistency, limited sex bias) (Couvry-Duchesne et al., 2017). In
212 addition, we have shown that the anxiety-depression score from the SPHERE
213 collected in adolescence predicts lifetime DSM-IV MDD diagnoses in young
214 adulthood in the BLTS sample.

215

216 Table 1: Detailed family structure of the QTIM and HCP samples

		QTIM	HCP
Total sample size (Individuals) for phenotypic analysis		834	890
Final sample size (individuals) for twin modeling		833*	853** ♣
Incl.	N complete trios	23	184
	Incl. N MZ pairs + one sibling	9	75
	N DZ pairs + one sibling	14	71
	N sibling trios	0	38
N complete pairs		275	109
	Incl. N MZ pairs	101	14
	N DZ pairs	142	19
	N twin-sib pairs	32	38
	N sibling pairs	0	38
N Singletons (214 individuals)		214	83
	Incl. N twins	181	33
	N non-twin siblings	33	50

217 * 1 individual from a family of 4 siblings was excluded for twin modelling (i.e. maximum family
218 size restricted to 3 individuals (twin and/or non-twin siblings))

219 ** 37 individuals from families with 4+ siblings were excluded with no effect on the sample
220 characteristics (mean age 28, sd=3.7, range 22-37, 56.5% females after exclusion)

221 ♣ To maximise the HCP sample 12 half-siblings were included (categorised as siblings in twin
222 analysis); this low number should not affect estimates from twin modelling.

223

224 Genetic parcellation of the cortex

225 Twin modelling contrasts MZ twins (who share the same genetic
226 information and familial environment) with DZ twins (who have the same
227 familial environment but share on average half of their genetic information) and
228 unrelated pairs (independent familial environment and genetic information). It
229 allows the estimation of the proportion of inter-individual differences

230 attributable to genetic variability in the population (narrow sense heritability or
231 additive genetic effects (A), the individuals' unique environment (E) and either
232 the familial environment (C) or genetic dominance (D) (Martin and Eaves, 1977;
233 Neale and Cardon, 1992; Verweij et al., 2012). Multivariate models can further
234 break down the sources of variances into common and trait specific, providing
235 an estimate of genetic and environmental correlations (Martin and Eaves, 1977;
236 Neale and Cardon, 1992).

237 Following prior work (Chen et al. 2013), and to reduce the number of
238 brain phenotypes tested (and the burden of multiple testing correction), we
239 estimated the matrices of genetic correlations between the vertex-wise SA and
240 CT measures (2305 left hemisphere, 2308 right hemisphere; medial wall vertices
241 excluded) within each hemisphere. To do this, we used bivariate AE models
242 (Neale and Cardon, 1992) using OpenMx (Boker et al., 2011), which implements
243 Full Information Maximum Likelihood (FIML) and allows some missingness in
244 the outcome variables. Omitting C/D terms in the models may result in a slight
245 overestimation of the genetic correlation, as only moderate shared
246 environmental sources of variance have been detected throughout the cortex
247 (Lenroot et al., 2009; Kremen et al., 2010; Eyler et al., 2011). We included a
248 maximum of 3 individuals (e.g. 1 twin pair and a sibling) per family. This
249 increases power to detect A and C/D compared to the standard twin design
250 (Posthuma and Boomsma, 2000).

251 Before estimating the genetic correlations, we residualised the CT and SA
252 measures to remove global effects (whole brain total SA, mean CT), as well as
253 effects of sex, age, acquisition orientation. We used a fuzzy clustering algorithm
254 to identify clusters of genetically correlated voxels (Kaufman and Rousseeuw,

255 1990; Chen et al., 2013) implemented in the R package “cluster” (R Development
256 Core Team, 2012; Maechler et al., 2016) and determined the optimal number of
257 clusters using silhouette coefficients (Kaufman and Rousseeuw, 1990; Chen et al.,
258 2013) (Figure 1). Such coefficients combine cluster cohesion (intra cluster
259 differences) and separation (inter cluster differences). A high coefficient
260 indicates better-separated clusters (Chen et al., 2013). The silhouette coefficients
261 plateaued after 10 to 14 clusters, for each hemisphere and measurement (Figure
262 1). We restricted our analyses to 12 clusters, which results in a relatively
263 parsimonious parcellation, to facilitate both comparisons across hemisphere and
264 with prior work (Chen et al., 2013). For each cluster, we calculated the mean
265 thickness and surface area, reducing the cortex to a total of 48 phenotypes. See
266 Figure 2 for cluster visualisation and labelling. See extended data 1-4 for
267 correspondence between cortical vertices and SA or CT clusters. In this analysis,
268 we preferred the clusters derived from our QTIM sample to those defined
269 previously in a sample of middle-aged male veterans from the US (Chen et al.,
270 2013). Though, it is interesting to observe that the clusters are largely conserved
271 across these two samples (Couvry-Duchesne, 2017).

272

273 Association between cortical measures and anxiety-depression

274 First, we tested the phenotypic association between the anxiety-
275 depression SPHERE score and both SA and CT for each of the genetic clusters. To
276 capture complex associations, we modelled linear, quadratic and cubic
277 relationships using natural splines, implemented in R (R Development Core
278 Team, 2012). Natural splines are piecewise polynomials (of 3rd order here) that
279 include constraints at the boundaries, forcing the tails of the association to be

280 linear, thus preventing diverging “unnatural” solutions. All models included the
281 3 splines orders, required to make the results interpretable. This approach can
282 detect both linear and U-shaped (quadratic) relationships, as well as changes
283 limited to the extreme of the anxiety-depression continuum (cubic), or even a
284 mixture of these. Compared to simple linear modelling, this should provide a
285 better fit to the data and may detect associations otherwise overlooked. In
286 addition, splines are more robust than polynomial regression as the splines
287 orders are constrained to be orthogonal, thus removing issues arising from
288 collinearity between the predictors. To limit the effect of outlier scores on the
289 results, we winsorised four observations with SPHERE sum scores greater than
290 20 (more than 4 SD away from the mean, in the heavy right tail).

291 To control for sample relatedness, we used a mixed model that integrates
292 the variance-covariance of observations via a 834x834 kinship matrix
293 (calculated with R package “kinship2” (Therneau and Sinnwell, 2015)). To
294 estimate the parameters of the models, we used the “hglm” package (Ronnegard
295 et al., 2010) that relies on extended quasi-likelihood (Lee and Nelder, 1996; Lee
296 and Lee, 2012; Ronnegard and Lee, 2013). The significance of the anxiety-
297 depression score (and all fixed effects) was tested using Students t-tests (test
298 statistic: $\beta/SD(\beta)$). We included as covariates the linear and non-linear
299 (quadratic and cubic) effects of age (at questionnaire and age at scan), sex, wave,
300 acquisition direction (coronal or sagittal), as well as mean vertex-wise SA/CT in
301 order to adjust for sex and age, which are in turn strongly associated with
302 head/body size/cortical surface area (sex, age) and thickness (age).

303 To avoid an overly stringent multiple testing correction, we estimated the
304 effective number of independent brain phenotypes from the eigenvalues of the

305 correlation matrix (Li and Ji, 2005). This method provides an accurate and fast
306 alternative to permutation tests (Li and Ji, 2005). We estimated the effective
307 number of genetic clusters to be 33 (out of a total of 48) (Li and Ji, 2005), after
308 regressing out the covariates (age, sex, acquisition, wave, mean vertex-wise
309 SA/CT) that tend to inflate the coefficients of the correlation matrix. This
310 translates into about 99 independent tests (as we are testing 3 effects: linear,
311 quadratic and cubic), which yields a significance threshold of $5.2E-4$ (Sidak,
312 1967) corresponding to a FWER (Family Wise Error Rate) less than or equal to
313 5%.

314 Then, we used bivariate (AE) twin models to estimate the genetic and
315 environmental correlations between cluster measurements and anxiety-
316 depression. Covariates used in the phenotypic association analysis were
317 regressed from the brain measurement and the residuals were used in the twin
318 analysis (see Table 1 for sample composition. In this paper, we only used
319 residuals in multivariate OpenMx models in order to reduce computational time
320 and directly included the covariates in all other models. We tested the
321 significance of the heritability estimates as well as the genetic and environmental
322 correlations using a likelihood ratio test on nested models (1 degree of freedom
323 tests). We corrected for the number of tests performed. Twin analyses were
324 performed in OpenMx (Boker et al., 2011). We also conducted a post-hoc
325 analysis to highlight which voxels from significantly associated genetic clusters
326 were driving the phenotypic association with anxiety-depression.

327

328 Post-hoc association analyses in the QTIM sample

329 In addition to the main analysis, we report the association of the SPHERE
330 scores with the surface and thickness measurements derived from the Desikan
331 anatomical atlas (Desikan et al., 2006), as well as a vertex-wise analysis. This
332 facilitates comparison with previous studies and shows that the reported
333 associations are not an artefact of our genetic parcellation. The same models and
334 covariates were used as for the main analysis. We estimated the significance
335 threshold to be $1.5E-4$ (108 effective phenotypes) for the associations with the
336 anatomical cortical regions. For the vertex-wise analysis, we used the brain-wide
337 significance threshold ($1.6E-5$, based on 10,000 effective phenotypes) calculated
338 by Medland et al., (Medland et al., 2014). Both significance thresholds took into
339 account that we tested 3 splines orders.

340

341 Replication Analysis

342 Participants

343 We used MRI and Achenbach Adult Self Report (ASR) questionnaire
344 (Achenbach, 2009) data for 890 participants (mean age 28 SD 3.7 range 22-37,
345 56.5% female) from the Human Connectome Project (HCP) (Van Essen et al.,
346 2012b; Van Essen et al., 2013) to replicate our results. In twin modelling, a
347 maximum of 3 participants per family was included, resulting in 184 complete
348 trios, 109 pairs and 83 singletons (see Table 1 for detailed breakdown).

349 We utilised pre-processed T1-weighted structural scans from the HCP
350 sample (Marcus et al., 2011). Minimal processing of the structural images by the
351 HCP team consists of removing spatial artifacts and T1 alignment (Van Essen et
352 al., 2012a; Glasser et al., 2013), using FSL (Jenkinson et al., 2002; Jenkinson et al.,
353 2012) and FreeSurfer (Fischl, 2012). We applied the same techniques used for

354 QTIM data to produce the vertex-wise and genetic cluster measures from the
355 pre-processed HCP surfaces.

356 We used the two ASR scores that were available: the anxiety-depression
357 syndrome based scale (Achenbach, 2009) and a DSM-oriented scale that we
358 constructed by summing the DSM-oriented scales for both anxiety and
359 depression (Achenbach et al., 2003). The correlation between the DSM-oriented
360 anxiety and depression scales was 0.66 (95%CI 0.62 0.70) and mostly driven by
361 common genetic sources of variance: $r_G=0.87$ (95%CI 0.70 1.00), $r_E=0.54$
362 (95%CI 0.42 0.64). These high phenotypic and genetic correlations suggest that
363 combining the 2 DSM-oriented scales is a valid approach. Participants reported
364 an average syndrome score of 5.7 (SD=5.2, range 0-33) and an average DSM-
365 oriented score of 8.0 (SD=5.6, range 0-35). The phenotypic correlation between
366 the DSM-oriented and syndrome anxiety-depression score was 0.90 (95%CI 0.89
367 0.91), with comparable genetic and environmental correlations ($r_G=0.91$, 95%CI
368 0.84 0.95, $r_E=0.89$, 95%CI 0.85 0.92). Notably, the genetic correlation between
369 the two scores was significantly different from 1 ($p\text{-value}=1.2E-4$), suggesting
370 that they may have some unique genetic sources of variance. All ASR scales have
371 previously shown good test-retest reliability and internal consistency
372 (Achenbach et al., 2005). In addition, the ASR syndrome based scale is heritable
373 through adolescence (Nivard et al., 2015) and captures a stable construct across
374 age and sex (Fonseca-Pedrero et al., 2012). DSM-oriented and syndrome based
375 scales appear to comparably predict affective disorder diagnoses (Najman et al.,
376 2008; Dingle et al., 2010; Dingle et al., 2011).

377

378 Replication Analysis

379 We considered for replication all significant associations between the
380 anxiety-depression (SPHERE) score and genetic clusters identified in the
381 exploratory analysis. As before, we tested for linear, quadratic and cubic
382 association with the brain phenotypes. We also corrected for sex, acquisition
383 variables, total SA or CT, linear and non-linear effects of age as well as familial
384 relatedness using a kinship matrix (Therneau and Sinnwell, 2015). We
385 controlled for multiple testing in the replication analysis by estimating the
386 effective number of independent phenotypes carried in the replication step (Li
387 and Ji, 2005). We further corrected for testing linear, quadratic and cubic
388 relationships and for considering the two ASR anxiety-depression scores. ASR
389 scores more than 4 standard deviations from the mean were winsorised to limit
390 the influence of outliers (heavy right tail) on the results.

391 We then explored the voxel-wise associations within replicating clusters
392 and further decomposed the associations into their genetic and environmental
393 components using bivariate AE twin models in OpenMx (Boker et al., 2011). As
394 before, we regressed out the covariates before twin modelling.

395

396 Code accessibility

397 The code/software described in the paper is freely available online at [retracted
398 for double-blind review]. The code is also accessible as extended data.

399

400 Results

401 Exploratory Analysis

402 Across the 48 brain clusters only one (cluster 6R for SA) comprising the
403 lower part of the occipital cortex was associated with the anxiety-depression

404 score from the SPHERE, after correcting for multiple tests (third order spline
405 association: $\beta=-0.037$, $SD=8.6E-3$, $rP=-0.10$, $p\text{-value}=2.4E-05$, Table 2). Surface
406 area in this cluster appeared to peak for participants with a median SPHERE
407 score (score = 2.8) and to be substantially decreased (up to 1 SD lower) in those
408 with a high-anxiety depression score (Figure 3a). Though, we should be cautious
409 interpreting the increasing surface area for participants with low SPHERE
410 scores, as it may be an artefact of plotting only the significant spline order.
411 Indeed, this increase is not observed when plotting the effect of all splines order
412 (dotted line, Figure 3a). We are confident the model is not over fitted as the 3
413 splines order only accounted for a 1.6% of the score variance and we did not
414 observe large standard errors for the estimates, which often indicate colinearity.

415 Fitting AE models, heritability was estimated to be 0.55 (95%CI: 0.32
416 0.73) for the 6R cluster for SA and 0.28 (95%CI: 0.14 0.41) for anxiety-
417 depression (SPHERE), but neither the genetic nor environmental correlations
418 between these traits reached significance (cubic correlations: $rG=-0.068$, $p\text{-}$
419 $value=0.35$ and $rE=-0.084$, $p\text{-value}=0.11$; linear correlations: $rG=-0.20$,
420 $pvalue=0.35$ and $rE=0.0091$, $p\text{-value}=0.91$). In comparison, heritability of the
421 other SA clusters ranged from 0.39 to 0.71 (0.32 to 0.63 for the CT clusters).

422 The 6R cluster for SA consists of a total of 94 vertices and is located in the
423 occipital cortex. It comprises most of the fusiform gyrus (56.6% or 43 vertices)
424 and the parahippocampal gyrus (56.5% or 13 vertices) as well as the lower part
425 of the lingual gyrus (33.8% of the gyrus or 22 vertices) and the medial part of the
426 lateral occipital cortex (12.6% or 12 vertices). To identify if one or more regions
427 of the 6R cluster were driving the cubic association between SA and the anxiety-
428 depression SPHERE score we restricted our analysis to the 94 voxels in the 6R

429 cluster. Using a significance threshold that accounts for the extra number of tests
430 performed ($p\text{-value} < 4.2E-4$), 61 voxels were found to be driving the non-linear
431 association between SA of the cluster 6R and the anxiety-depression SPHERE
432 score (p-value range: $5.9E-06$ $3.8E-04$, correlation range: -0.13 -0.091 , Figure
433 4a). When decomposing the phenotypic association using an AE model, none of
434 the genetic or environmental correlations was significantly different from zero.

435

436 Table 2: Summary of discovery (QTIM sample) and replication (HCP sample) analysis

		First order spline (linear)			Second order spline (quadratic)			Third order spline (cubic)		
		<i>Beta</i> <i>r</i>	<i>(se)</i>	<i>p-value</i>	<i>Beta</i> <i>r</i>	<i>(se)</i>	<i>p-value</i>	<i>Beta</i> <i>r</i>	<i>(se)</i>	<i>p-value</i>
QTIM	SPHERE scale	0.010 r=0.030	(6.0E-3)	0.094	-0.021 r=-0.069	(6.5E-3)	1.6E-3	-0.037 r=-0.10	(8.6E-3)	2.4E-5
	DSM- oriented scale	-0.18 r=-0.040	(0.060)	3.2E-3	0.015 r=0.0022	(0.12)	0.89	0.15 r=0.022	(0.120)	0.13
HCP	ASR syndrome based scale	-0.032 r=-0.0064	(0.076)	0.68	-0.044 r=-0.0088	(0.10)	0.67	0.079 r=0.012	(0.15)	0.61

437 Significant associations (after multiple testing correction) are reported in bold.

438

439 Replication Analysis

440 We found a significant (first order spline) association between the DSM-
441 oriented score for anxiety-depression and SA of the 6R cluster (Table 2), using a
442 significance threshold of $8.5E-3$, which corresponds to a $FWER < 5\%$ considering
443 6 independent tests (3 tests of association for 2 anxiety-depression scores). This
444 association supports our finding in the QTIM data set, of a reduction of SA ($r = -$
445 0.040 , $p\text{-value} = 3.2E-3$) in participants with moderate anxiety-depression score
446 (25-95 percentile) (Figure 3b), which we reported above.

447 In the HCP dataset, SA for the 6R cluster was moderately heritable (0.63 ;
448 $95\%CI$ $0.52 - 0.72$), after regressing mean SA and the other covariates, as well as
449 the anxiety-depression scores: 0.36 for the DSM oriented ASR ($95\%CI$: $0.20 -$
450 0.51) and 0.41 for the syndrome based ASR ($95\%CI$ $0.24 - 0.56$). When breaking
451 down the phenotypic association between DSM-oriented scale and 6R SA, neither
452 genetic nor environmental components reached significance (linear correlations:
453 $rG = -0.090$, $p\text{-value} = 0.37$ and $rE = -0.032$, $p\text{-value} = 0.59$). Likewise, when using the
454 ASR syndrome based scale ($rG = -0.078$, $p\text{-value} = 0.56$ and $rE = -0.059$, $p\text{-}$
455 $value = 0.49$). Heritability of other SA clusters ranged from 0.46 to 0.82 in the HCP
456 (0.42 to 0.72 for CT clusters).

457 We then tested the association between each voxel of the cluster with the
458 DSM-oriented scale. We estimated the number of effective independent voxels to
459 be 7 (Li and Ji, 2005) after regressing out the covariates' effect on the vertex-
460 wise measurements. Thus, we used a significance threshold of $3.6E-3$, to account
461 for 13 independent tests (6 previously and 7 for the voxel-wise testing).
462 Nineteen (out of 94) voxels located in the medial posterior part of the cluster
463 survived multiple testing correction ($-\log(p\text{-value}) > 2.4$, r in $-0.040 - 0.068$, Figure

464 4b). Thirteen of these were located in the lingual gyrus and were also associated
 465 with the anxiety-depression scores in QTIM. A further 3 neighbouring vertices
 466 from the fusiform gyrus were also consistently associated to anxiety-depression
 467 across the two samples (Table 3). As previously, vertex-wise genetic and
 468 environmental correlations did not reach significance.

469

470 Table 3: Summary of post-hoc vertex-wise association testing for SA of the 6R cluster

	N total vertices in each gyrus	N vertices for cluster 6R		N vertices associated in vertex-wise analysis- QTIM	N vertices associated in vertex-wise replication analysis - HCP
		N	MNI coordinates		
Parahippocampal	23	13	x in 14 33 y in -64 -50 z in -19 -6	13	0
Lingual	65	22	x in 7 32 y in -107 -68 z in -4 6	22	13 #
Fusiform	76	43	x in 24 43 y in -102 -58 z in -19 3	22	3•
Lateral Occipital	95	12	x in 21 45 y in -109 -91 z in -8 0	0	2

471 # All vertices associated in QTIM. Vertices correspond to numbers 144, 570, 573, 2220, 2244,
 472 2252, 2253, 2258, 2262, 2264, 2265, 2266, 2269 of the FreeSurfer fsaverage4 parcellation (MNI
 473 coordinates: x in 7 20, y in -107 -82 and z in -3 6).

474 • All vertices associated in QTIM. Vertices correspond to number 2263, 2267 and 2270 of the
 475 FreeSurfer fsaverage4 parcellation (MNI coordinates: x in 24 27, y in -98 -92, z in 0 3).

476

477 Post-hoc association analyses

478 The association analysis between the SPHERE scores and anatomical regions
479 from the Desikan atlas returned one significant association with surface area of
480 the right lingual gyrus ($p=9.5E-5$). As per the main QTIM analysis, the association
481 was negative and with the 3rd order spline ($\beta=-330.7$, $SD=83.9$, $r=-0.13$). No
482 significant association was found at a vertex-level for surface area or cortical
483 thickness.

484

485 Discussion

486 Here, we used twin modelling to derive a genetic parcellation of two
487 characteristics of the cortical ribbon (i.e. surface area and thickness), reducing
488 each hemisphere to 12 measurements. The surface area of one region (cluster
489 6R) comprising part of the lingual, fusiform, parahippocampal and lateral
490 occipital gyri showed a significant association with the anxiety-depression score
491 from the SPHERE in the QTIM sample. We partially replicated this association in
492 an independent sample of similar age, using data from the HCP, in which we
493 found an association between the DSM-oriented score and surface area in the 6R
494 cluster. Both associations suggested lower surface area in participants with
495 anxiety-depression scores above average (50-95% quantiles). A post-hoc vertex-
496 wise analysis suggested that the lingual vertices (and to a lesser extent fusiform)
497 might drive the association. We cannot conclude whether the observed
498 association is specific to the right hemisphere due to differences in the cortical
499 parcellation between hemispheres (Figure 2, in the left hemisphere the occipital
500 cortex is composed of a unique cluster).

501 Interestingly, a meta-analysis (472 cases, 680 controls, from 12 published
502 studies, mean age 43, partially medicated) identified significant reductions of GM

503 density (Ashburner and Friston, 2000) (often considered as a proxy for volume)
504 in right lingual, fusiform and parahippocampal gyri (Arnone et al., 2016). The
505 authors found little evidence for publication bias and the voxel-based
506 morphometry approach cannot be confounded by head/body size. This echoes
507 results from the meta-analysis by the ENIGMA-MDD group, which reported a
508 significant reduction of right lingual gyrus surface area in adolescent depression
509 (237 cases, 294 controls) (Schmaal et al., 2016) (non-significant association for
510 fusiform or parahippocampal). Though the meta-analysis by ENIGMA did not
511 correct for total SA and included medicated cases, it is also not totally
512 independent from the present study as it included QTIM participants (26 cases,
513 140 controls). Even so, in other work, right lingual GM density was identified as
514 a predictor of anti-depressant response, with higher density predicting better
515 response to first treatment (Jung et al., 2014). The non-respondent group
516 showed lower performance in 3 aspects of the Stroop test (neutral word, colour
517 word and error control), and a post-hoc analysis in cases showed right lingual
518 density to be associated with better error control (Stroop), and non-verbal
519 memory (Rey-Kim memory test). This is coherent with neurological case reports
520 that highlight the crucial role of the lingual gyrus in visual memory
521 (Bogousslavsky et al., 1987). Furthermore, impaired visual memory has been
522 associated with 1st episode MDD (Lee et al., 2012) and with adolescent MDD
523 (Baune et al., 2014); lower Stroop accuracy has also been reported in adult MDD
524 (Snyder, 2013), while more research is needed in pediatric depression (Vilgis et
525 al., 2015). Finally, reduced right lingual volume was also reported in adult MDD
526 cases (partially medicated) (Lee et al., 2011), following traumatic brain injury
527 (Maller et al., 2014).

528 The reported right lingual gyrus volume (or density) reduction in
529 depression may be driven by reduced surface area, at least in early adulthood,
530 which may be associated with visual memory and attention deficits in
531 depression. In addition, the right lingual surface area may be associated with
532 anti-depressant response and cognitive functions in MDD patients (Jung et al.,
533 2014), which might explain why the findings are not consistent across samples
534 that may differ in term of medication status, response to treatment or cognitive
535 abilities. For example, screening in QTIM may exclude individuals with extreme
536 levels of anxiety-depression, and we may be only observing part of the
537 association.

538 Reduction of right fusiform gyrus volume has also been reported in off-
539 medication adult patients (Lener et al., 2016), but several large studies of older
540 (medicated) participants attributed this reduction to reduced cortical thickness
541 (Canu et al., 2015; Schmaal et al., 2016). In addition, cortical thickness of the
542 right fusiform in MDD cases with comorbid generalised anxiety was even more
543 reduced (Canu et al., 2015). The right fusiform gyrus has a central role in face
544 perception (Haxby et al., 2000), in neurological case reports of prosopagnosia
545 (face blindness) (Whiteley and Warrington, 1977; Damasio et al., 1990; Derenzi
546 et al., 1994), behavioural studies (Rhodes, 1993) and imaging studies (McCarthy
547 et al., 1997; Haxby et al., 2000). Lower right fusiform grey matter density was
548 observed in developmental prosopagnosia (Garrido et al., 2009) and in
549 congenital prosopagnosia (Behrmann et al., 2007). In addition, both studies
550 reported an association between right fusiform volume and performance in face
551 identification (Behrmann et al., 2007; Garrido et al., 2009). Reviews on adult
552 depression suggest a generally reduced accuracy of face expression evaluation,

553 coupled to increased attention and response bias in expression evaluation
554 toward sadness (Bourke et al., 2010).

555 In adolescent depression, more research is needed to confirm the mixed
556 evidence of impairment in face (and face expression) processing (Vilgis et al.,
557 2015). Lower fusiform gyrus surface area in young adults might relate to the
558 poorer face recognition and processing of face expressions reported in adult
559 depression but we did not have the cognition data to test such hypothesis.
560 However, our results only weakly point to the fusiform gyrus and do not align
561 with prior studies that only reported a reduced fusiform thickness associated
562 with MDD (Canu et al., 2015; Schmaal et al., 2016). Overall, we cannot rule out
563 that medication or sample age might explain some of these differences, as well as
564 the hypothesis of delayed cortical development in depressed participants,
565 implying a normalisation of cortical surface in mid-adulthood that leaves a more
566 permanent decrease in thickness (Schmaal et al., 2016).

567 In line with prior publications (Panizzon et al., 2009; Chen et al., 2013)
568 that reported significant heritability of cortical thickness and surface, we
569 confirmed that the SA of the 6R cluster is significantly heritable ($h^2_{QTIM}=0.55$,
570 $h^2_{HCP}=0.63$), even after correcting for total SA, suggesting that this region is
571 influenced by specific genetic sources of variance, that do not contribute to
572 global SA. In addition, all anxiety-depression scores were comparably heritable
573 ($h^2_{SPHERE}=0.28$, $h^2_{DSM-ASR}=0.36$, $h^2_{ASR}=0.41$), in line with the known heritability of
574 depression from twin and family studies (Sullivan et al., 2000; Kendler et al.,
575 2006; Polderman et al., 2015). In the bivariate analysis, because of the modest
576 phenotypic association, we had limited statistical power to determine whether
577 common genetic or environmental factors drove the observed phenotypic

578 association. Indeed, a post hoc power analysis indicates that we had, at best, 10%
579 power to detect the observed genetic correlation in QTIM and HCP (taking into
580 account observed heritability and effect sizes, assuming an AE model and a risk
581 $\alpha=5\%$). A combined analysis of the two samples would not confer a statistical
582 power greater than 22%. The limited power may explain the instability of the
583 correlations estimated in both samples, even if differences in anxiety-depression
584 questionnaire, sample composition or recruitment could also be at play.

585 The reported heritabilities of the scores are consistent with the diathesis-
586 stress model of anxiety-depression, in that genetic liability and environmental
587 factors both contribute to the risk of depression (Rosenthal, 1963), sometimes in
588 a multiplicative manner (Peyrot et al., 2014; Musliner et al., 2015; Mullins et al.,
589 2016; Colodro-Conde et al., 2017). More generally, the identification of brain
590 endophenotypes should lead to refine this model by providing insight into the
591 brain networks and behavioural mechanisms that are influenced by the genetic
592 and environmental risk factors. We still lack evidence to provide a credible
593 interpretation for the association reported here, which may require reconciling
594 evidence produced from complementary types of imaging. For example, reduced
595 glucose metabolism in the right lingual gyrus has been linked to sleep anomalies
596 in depression (Germain et al., 2004; Fitzgerald et al., 2008) but more work is
597 need to investigate whether structural cortical markers may tag the same brain
598 pathways and contribute to the same symptoms.

599 Our study has several limitations. Firstly, we did not perform a strict
600 replication as QTIM and HCP differed in the anxiety-depression scores used, time
601 difference between scoring and scanning, and shape of the relationship (3rd vs.
602 1st order splines). Thus, differences in studies design (different timeframes,

603 medication, population genetics and environment, ascertainment bias...) might
604 be partially responsible for the different effect size observed in QTIM and HCP. In
605 addition, the SPHERE and ASR both aim to measure anxiety-depression but they
606 are composed of different questions and may be sensitive to different ranges of
607 symptoms or different stages of the disorder (e.g. subclinical for the SPHERE and
608 clinical for the DSM based ASR). Furthermore, the association between 6R SA
609 and anxiety-depression scores in the HCP was only observed using the DSM-
610 oriented score, which we attribute to differences in scores distribution making it
611 harder to detect non-linear associations. Indeed, despite high (phenotypic and
612 genetic) correlations between the scores ($r=0.90$, $r_G=0.91$), the correlation
613 between spline orders was lower ($r=0.79$ between first order splines of the 2
614 anxiety-depression scores, 0.56 between second order and 0.73 between third
615 order). This may reflect differences of scale and distribution. Differences in
616 genetic sources of variances (genetic correlation different from 1) are unlikely to
617 completely explain the differences in results, as the correlation remained high.
618 Nonetheless the different results between QTIM and HCP as well as the
619 inconsistent associations in the HCP sample using different scores remain
620 important limitations of this study. These limitations call for further
621 investigation of the relationship between lingual gyrus and anxiety-depression.

622 The small effect size of the association in QTIM and HCP can also be seen
623 as a limitation, as lingual gyrus SA does not explain much of the anxiety-
624 depression variability and would have a very limited predictive power. However,
625 our result is in line with the effect sizes observed in larger case-control studies
626 that suggest that no single brain marker of depression explains much of the
627 disorder (Schmaal et al., 2016). A parallel to this is the small SNP effects

628 observed/found in GWAS of complex traits (Visscher et al., 2017). Thus, despite
629 the small association between cortical regions and the trait/disease, their impact
630 on brain networks and functions may be much larger. That is, a small association
631 does not necessarily preclude a biological or cognitive relevance.

632 Further limitations include our clustering approach, which may make it
633 harder to detect localised structural changes associated with anxiety-depression,
634 as the thickness and surface are averaged over large cortical regions. In addition,
635 we did not investigate sex specific changes that may reflect depression subtypes
636 with different aetiology (Kendler et al., 2001; Kendler and Gardner, 2014). As per
637 the hypothesis of age specific markers reported by the ENIGMA-MDD (Schmaal
638 et al., 2016), age was not significant in our analysis, which prevented us from
639 testing for an age interaction with anxiety-depression scores. However larger
640 samples would be required to overcome the multiple testing correction burden
641 from a vertex-wise analysis or the loss of power resulting from study
642 stratification. Finally, the interpretation of our results is limited by the relative
643 lack of robust research on cognitive and imaging aspects of pediatric depression
644 and of normal neuro-development. The Research Domain Criteria (RDoC) may
645 help in making connections between cognition, imaging, genetics and psychiatric
646 illnesses.

647 In summary, we proposed a candidate endophenotype for depression,
648 which is heritable, and shows a replicable phenotypic association with anxiety-
649 depression scores. The vertex-wise post-hoc analysis suggested that a reduction
650 of SA in the ventral part of the lingual gyrus could drive the observed association.
651 Longitudinal studies beginning in adolescence are emerging (Luby et al., 2016;
652 Schmaal et al., 2017) and will help clarify the temporal and causal relationships

653 between brain development, cognition and mental health.

654

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983 Figure 1: Silhouette coefficients of the clustering scenarios

984 Vertical dashed line corresponds to 12 clusters per hemisphere and
985 measurement, which we used to parcellate the cortex.

986

987 Figure 2: Summary of genetic parcellation of the cortical thickness and surface
988 area

989 Clusters are labelled 1L to 12L for the left hemisphere and 1R to 12R for the right
990 one. Extended Data Figure 2-1, Figure 2-2, Figure 2-3, and Figure 2-4 describe
991 the genetic clusters used in this analysis.

992

993 Figure 3: Non-linear associations between anxiety-depression and surface area
994 of the genetic cluster 6R

995 The solid line represents the regression effect of the significant spline order. The
996 dashed line combines the association for the linear, quadratic and cubic splines
997 and the anxiety-depression SPHERE score and the SA of the 6R cluster. The y-
998 axis corresponds to SA after removing the effect of the intercept and all other
999 covariates. The vertical dashed bars indicate the 5, 25, 50, 75 and 95% quantiles
1000 of the scores distributions

1001 a) Association between SA and cluster 6R SA and the SPHERE anxiety-depression
1002 score (QTIM sample). The reduction of surface area is observed in participants
1003 with SPHERE score greater than 2.9 (median score, see solid line: 3rd order
1004 natural spline). We interpret the increase of the (solid) regression line for SPHERE
1005 score below median as an artefact of plotting the significant spline order only. This
1006 effect is not observed when the 3 splines order are combined

1007 b) Association between SA for cluster 6R and the DSM-oriented (left panel) and
1008 ASR syndrome based (right panel) anxiety-depression scores. A consistent
1009 reduction of surface area was observed in participants with ASR DSM-oriented
1010 scores between 4 and 19 (25% - 95% centiles; see solid line: 1st order natural
1011 spline). No significant association was found with the ASR syndrome based scale.
1012

1013 Figure 4: Vertex-wise phenotypic association of surface area with anxiety-
1014 depression scores within cluster 6R

1015 Bottom view (top) and medial view (bottom) of the right hemisphere. The left
1016 panels show the cubic effect sizes (correlations) for each vertex of the 6R cluster
1017 for SA (94 vertex). The right panels show the significance (p-value) of the
1018 association in $-\log_{10}$ scale. Vertices represent the intersections of the triangular
1019 mesh reconstructed by FreeSurfer to model cortical surfaces.

1020 a) Vertex-wise association between SA and anxiety-depression (SPHERE score)
1021 in the QTIM sample (N=833). The significance threshold of $4.2E-4$ is reached for
1022 p-values greater than 3.4 in the log scale.

1023 b) Vertex-wise association between SA and anxiety-depression (DSM-oriented
1024 score) in the HCP sample (N= 890).

1025

1026 Extended data

1027 Figure 2-1: Description of vertices in each of left CT clusters.

1028 Figure 2-2: Description of vertices in each of left SA clusters.

1029 Figure 2-3: Description of vertices in each of right CT clusters.

1030 Figure 2-4: Description of vertices in each of right SA clusters.

1031

1032 Extended Data 1: R code used for all the analyses







